

# Are Retirement Planning Tools Substitutes or Complements to Financial Capability?\*

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August 17, 2022

## Abstract

We conducted a randomized controlled trial to understand how a web-based retirement saving calculator affects workers' retirement-savings decisions. In both conditions, the calculator projected workers' retirement income goal. In the treatment condition, it additionally projected retirement income based on defined-contribution savings, prominently displayed the gap between projected goal and actual retirement income, and allowed users to interactively explore how alternative, future contribution choices would affect the gap. The treatment increased average annual retirement contributions by \$174 (2.3%). However, effects were larger for those with greater financial knowledge, suggesting this type of tool complements, rather than substitutes for, underlying financial capability. *Word Count: 100 words*

**Keywords:** retirement planning, retirement saving, exponential-growth bias, present bias, fi-

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\*Acknowledgments: The pre-analysis plan for this project was registered at the [Social Science Registry AEARCTR-0002129](#). Human subjects protocol was determined as exempt by Stanford IRB (#38948), University of Minnesota IRB (#1607E90001) and Claremont Graduate University IRB (#2813). This research was supported by the U.S. Social Security Administration through grant #5-RRC08098400-10 to the National Bureau of Economic Research as part of the SSA Retirement Research Consortium and the Laura and John Arnold Foundation through a grant to the Stanford Institute for Economic Policy Research (SIEPR). The findings and conclusions expressed are solely those of the authors and do not represent the views of SSA, any agency of the Federal Government, the NBER, the Laura and John Arnold Foundation, SIEPR or any other institution with which the authors are affiliated. We are grateful to Paula Gablenz and Konhee Chang for exceptional research assistance. Declaration of interests: none. ©2018 Goda, Levy, Manchester, Sojourner, Tasoff, and Xiao. All rights reserved.

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nancial literacy, financial capability

# 1 Introduction

Determining how much to save for retirement is a complex problem that, in the era of defined contribution (DC) retirement saving plans, largely falls on the individual. Solving for one’s optimal retirement saving contribution in a given year requires simultaneously setting a target income in retirement and determining what contribution path enables one to meet that goal, taking into account investment returns, expected retirement age, and other sources of retirement income, such as Social Security and defined benefit pension income. Navigating this problem is challenging due to its high dimensionality, considerable uncertainty, and the limited opportunity to learn from mistakes.

There is reason to believe that many are not well-equipped to solve this complex problem. Rates of understanding for basic financial concepts are low ([Lusardi and Mitchell, 2014](#)). Further, limited financial understanding is one explanation for the disproportionate influence of defaults – which dictate employee outcomes when no choice is made – on participation and contribution decisions ([Madrian and Shea, 2001](#); [Beshears et al., 2009](#)). Individuals may look to default settings as implicit saving advice, yet such settings may not be aligned with individuals’ retirement lifestyle goals. In addition, [Goda et al. \(2019\)](#) show that low financial literacy and lack of understanding of exponential-growth bias are associated with lower retirement wealth accumulation among retirement-age individuals. These findings underlie public policy concern regarding the extent to which low financial capability fuels the limited retirement savings observed among many individuals.

Plan sponsors and academic researchers have sought to improve retirement saving decisions by supporting employee decision-making through information campaigns and via online retirement saving tools or calculators. Three key factors that determine whether these informational interventions are likely to be successful in addressing inadequate retirement saving are: 1) who selects into using them, 2) how they affect contribution behavior among those who use the tool, and 3) how the intervention differentially affects financially more vulnerable populations. Are such tools effective at raising financial decision-making capacity across the board, including those with limited financial literacy, or do the tools themselves require a sufficient understanding of financial concepts in order to be effective? That is, are such tools a substitute for existing financial knowledge or

a complement? Often plan sponsors introduce decision-support tools with the goal of increasing participation among those with lower financial knowledge, and thus implicitly assume these tools are substitutes.

To address these questions, we conduct a randomized controlled trial among employees at the U.S. Office of Personnel Management (OPM), an agency of the federal government. Federal employees have an employer-sponsored retirement savings program similar to a 401(k), called the Thrift Savings Plan (TSP), in which agencies match employee contributions. Our design randomly assigns employees to receive one of two online retirement saving tools: a treatment and an active control. Both tools elicit information on the participant’s desired lifestyle in retirement, current earnings, and expected retirement age in order to display their target retirement income as well as collect information on inputs to a retirement income projection. The tools differ in how complete this projected income calculation is. The “treatment tool” incorporates expected social security, federal defined-benefit pension income, and existing TSP savings and contribution levels into the projection, allowing participants to see whether the projection aligns with their target and dynamically assess how their TSP contributions map to their retirement income. In contrast, the “active control tool” omits retirement income stemming from TSP in the income projection. Instead, participants in the active control are asked to make their own assessment as to how much additional retirement income their accumulated TSP savings and future contributions will provide to assess whether they are on track to meet their retirement income target.

The difference between these two conditions isolates the effect of computations that map current behavior to financial resources in retirement. The accuracy of this mapping may vary as a function of an employee’s financial capability and willingness to engage in effortful thinking and planning. Past research has specifically implicated exponential-growth bias, present bias, and financial illiteracy as attributes implicated in low retirement savings (Goda et al., 2014; Brown and Previtro, 2014; Goda et al., 2019; Lusardi and Mitchell, 2011a). The additional information provided in the treatment removes the need to make exponential computations which require effort that is prone to postponement by present-biased individuals. Therefore, our treatment is designed specifically to overcome the exponential-growth bias and present bias that would otherwise lead to suboptimal decision-making.

First, we find that approximately half of employees (48 percent) select into using the tool, and

selection is correlated with pre-intervention TSP contributions. Next, we evaluate whether the treatment tool affected TSP contributions relative to the active control, and how the effect varied across employees using a treatment-on-the-treated (TOT) estimation approach. Overall, we find that the treatment increased average annual contributions by \$174 among those who used it relative to those using the control tool. We examine heterogeneous treatment effects across multiple measures of financial capabilities. We find that the treatment effect is significantly greater for those with higher measures of financial literacy, a college degree, and a higher financial-capability index score derived from factor analysis. We do not find evidence that exponential-growth bias, present bias, pre-intervention contributions, or other factors derived from factor analysis significantly predict the treatment effect.

Our study makes two main contributions to the literature. First, we find evidence that online retirement-income tools, specifically the part of the treatment focused on exponential computations, lead to modest increases in retirement savings. These findings are similar to [Goda, Manchester and Sojourner \(2014\)](#), who randomize a mailed leaflet to university employees on retirement savings, but differ in that the treatment condition involves a greater amount of engagement than a passive mailing. Second, we find evidence that retirement-income projections are complements rather than substitutes to financial capability. The online tool delivery provides the ability to track tool users and link engagement with the tool to outcomes. Selection into treatment is higher among those with higher pre-intervention contributions, and the treatment effect is larger for those with higher financial literacy and education. This finding is important as it suggests that retirement planning tools are unlikely to be sufficient to overcome biases that prevent optimal decision-making.

## 2 Experimental Design and Data

### 2.1 Retirement Plan Setting

As federal workers, OPM employees participate in a DC plan known as Thrift Savings Plan (TSP), in which the employer makes a base contribution of 1 percent of pay and matches employee contributions up to 5 percent of pay.<sup>1</sup> Employees can contribute up to the IRS maximum each year,

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<sup>1</sup>The agency matches dollar-for-dollar on an employee’s contributions up to 3 percent of pay and \$0.50 to the dollar for the next two percent of pay.

which was \$18,000 in 2017. Employees are also covered by a defined benefit pension.<sup>2</sup> Employees may elect to invest their contributions in five different funds or a lifecycle option, which is a mix of the other funds based on the employee’s age.

A 2015 TSP report indicated that approximately half of federal employees were not contributing enough to TSP to maximize the agency match OPM (2015). The proportion qualifying for the full match is even lower for recent hires, who are covered by a 3 percent automatic enrollment provision introduced in 2010. Concern about employees failing to maximize the match and the influence of automatic enrollment on contribution rates left OPM leaders seeking to develop an effective online retirement saving tool to improve TSP contribution decisions for federal employees.

## 2.2 Intervention

In partnership with OPM, we designed both a treatment and an active control version of the new online retirement saving tool with the aim of: 1) providing employees with both a target retirement income and a projected retirement income, and 2) isolating the effect of translating their TSP asset level and any potential contribution stream into a projected retirement-income stream on outcomes. The tool rolled out in November 2017. The two versions of the tool – treatment and active control – were made as similar as possible except that the active control did not provide any information on how TSP balances and contributions translated into retirement income. This allows us to isolate the effect of the income projection from any other tool features. The tool begins by asking the user a series of questions to determine their target income in retirement, such as their date of birth, when they started working for the federal government, their annual salary, their expected retirement age, and their desired lifestyle in retirement: 70, 85, 100, or 115 percent of current pay. Participants are visually shown the “goal” as a vertical bar, represented as a monthly annuitized income target for retirement. Then participants are asked questions to produce a projected retirement income based on their current assets and saving rate, including their TSP account balance, TSP annual contribution, pension coverage, and Social Security expectations. The main difference between the treatment and active control conditions is that the income projection of the former uses all information provided, while in the active control, it provides projections only based on pension and

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<sup>2</sup>Employees hired before 1984 are covered by a more comprehensive DB plan and receive no base and no match on employee contributions to TSP, yet can contribute on their own up to the individual maximum.

Social Security income, and states that retirement income from TSP is an additional amount on top of these other sources. Figure 1 shows the difference between the treatment and active control conditions in terms of the core visual that compares the retirement income goal and the retirement income projection. Screenshots of the entire tool are available in [Appendix E](#).

After displaying the projection, the treatment tool allows users to use sliders to adjust TSP contributions to see how the projection changes relative to the goal in real time. Some parameter values of the economic environment are needed to create the income projection such as the income growth, inflation rate, and expected real rate of return. The tool calculation formula is presented in [Appendix F](#). We provide default values for annual income growth and for inflation of 3 percent and 2 percent, respectively, which can be modified by the user. Because the real rate of return depends on one’s retirement portfolio, there may be considerable variation in people’s expected rate of return and desired lifestyles for retirement. We randomize these assumptions to test whether these default parameters affect saving behavior. The default rate of return is randomly assigned to 5 percent or 8 percent, and desired retirement lifestyle is randomized to 85 percent or 100 percent of income.<sup>3</sup> As with the other assumptions, these parameters could be modified by the user.

Both versions of the tool end by showing participants a printable summary of their current TSP contribution levels and a link to the TSP website and phone number with instructions on how to update contribution rates. The printable summary for the treatment tool also includes the last slider position for the TSP election.

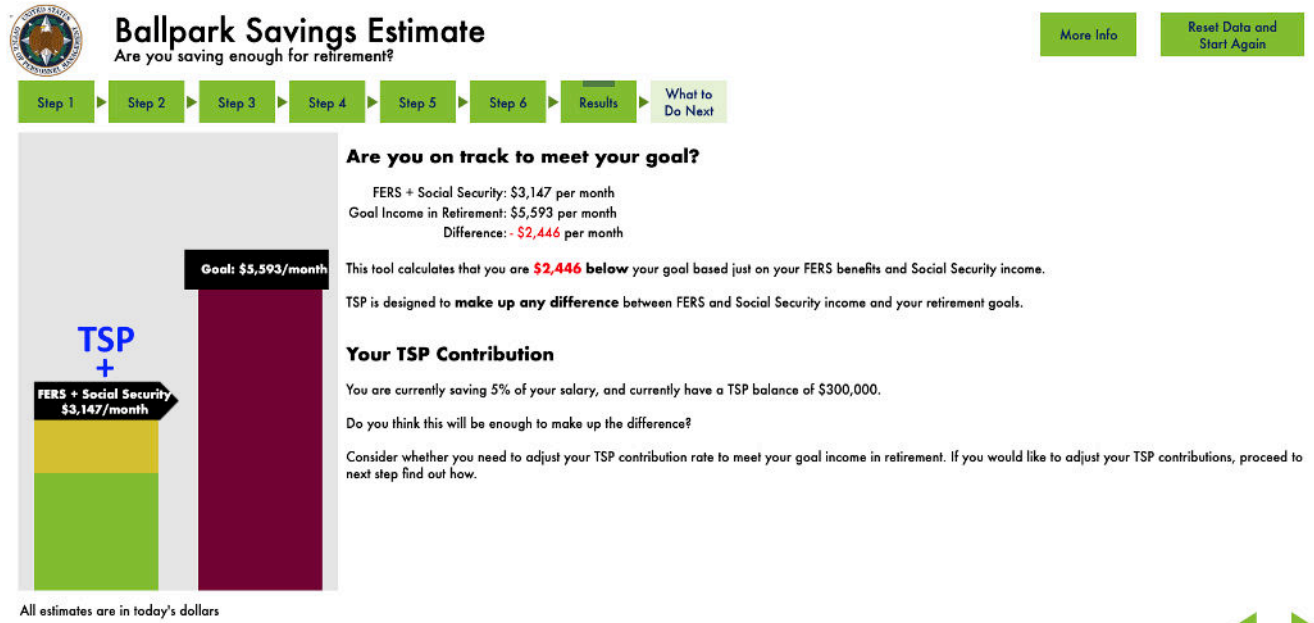
Prior to the intervention, we surveyed the employees to acquire background characteristics and elicit behavioral parameters that are not present in administrative data. The survey was fielded between March 29, 2017 and April 14, 2017. OPM emailed each employee an initial survey invitation and two reminders to non-respondents. Of the 5,426 employees, 1,435 completed the survey, a 26 percent completion rate. Through the survey, we measure financial capabilities, including exponential-growth bias (EGB), financial literacy, and college degree completion. The survey also elicited time preferences, including the long-term discount rate and a measure of present-biased preferences.

We randomly assigned the 5,426 unique individuals employed at OPM in December 1, 2017

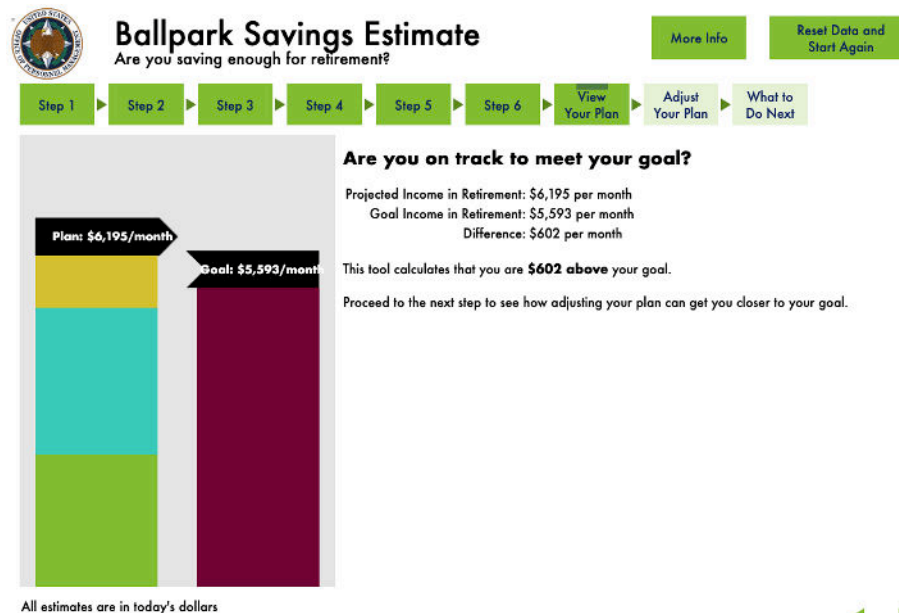
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<sup>3</sup>The results of these regressions are available in [Appendix B](#) in Tables B.11–B.13. The high values of these default parameters had an insignificant effect on savings relative to the low values. Participants have the option to change these assumptions using sliders and can view how they change the income projection.

Figure 1: Screenshot of Examples from Each Treatment Condition



(a) Active Control



(b) Treatment



to have access with equal probability either to the treatment tool or the active control tool. We stratified participants based on survey response with 50 percent of responders and 50 percent of non-responders each getting the treatment condition. Within a survey-response group (completers/non-completers) we stratified on total pay, age, TSP total amount, and gender. Survey completers were also stratified on their mean response to the EGB elicitation and mean response to the time-preference elicitation. OPM emailed each employee a personalized link to the appropriate version of the tool. Employees received an invitation to use the tool on December 1st, 2017. Subsequent reminder emails were sent to those who had not yet clicked the link on December 7th, December 18th, and January 11th. There was no differentiation in the invitation emails between the treatment and active control groups.

## 2.3 Data and Analysis Samples

Our data include each individual’s monthly TSP contribution elections and demographic characteristics from administrative HR records from August 2014 to April 2018. We match these data with survey data collected in March and April of 2017, and data on whether each individual chose to use the assigned tool or not: 2,625 (48 percent) did.

We use two analysis samples. The first just relies on administrative records including TSP contributions and employee characteristics recorded in HR files. This sample consists of the 2,625 unique employees who used the tool and their 152,198 total individual-by-month observations. The second analysis combines the TSP and HR records with survey responses and captures 1,435 unique individuals with 85,974 total individual-by-month observations. [Appendix C](#) presents a schematic of these samples.

We examine whether there were significant differences in baseline characteristics between individuals assigned to active control versus treatment ([Appendix Table A.1](#)). The joint test of null difference across baseline characteristics has a  $p$ -value of 0.96, reflecting successful random assignment. [Appendix Table A.2](#) compares survey completers and non-completers on administrative variables, which are fully observed. Survey completers are older, whiter, higher-paid, and contribute more than non-completers. To clarify which characteristics are most strongly associated with response conditional on the other characteristics, [Appendix Table A.3](#) reports estimates from a logit model of survey response. In this model, many observable characteristics predict response

but age, pay, and length of tenure do not.

## 2.4 Survey Measures

We perform our primary heterogeneity analysis on the subsample who completed the survey. Below we describe our measures of financial capability, which are central for assessing heterogeneous treatment effects, and the elicitation of time preferences. Finally, we present findings from an exploratory factor analysis of covariate space that shows the construction of a factor that aligns with financial capability.

### 2.4.1 Exponential-growth bias

We hypothesize that exponential-growth bias plays an important role in creating a gap between individuals' ideal savings rate and their actual savings rate. Exponential-growth bias is the tendency to neglect compound interest (Stango and Zinman, 2009). Forecasting one's retirement savings without the use of a tool requires considerable sophistication. The lack of an accurate forecast along with exponential-growth bias may cause people to underestimate the benefits of saving for retirement. Because the intervention operates by explicitly computing the exponential growth of the user's savings (along with other computations), those with greater bias may benefit from the intervention more. More precisely, because undersaving is likely a larger problem than oversaving (see for example Goda et al., 2019, who show that exponential-growth bias is correlated with lower retirement savings), people with more exponential-growth bias may exhibit larger treatment effects.

We use the parametric model of Levy and Tasoff (2016) given below.

$$p(\vec{r}, t; \alpha_i) = \prod_{s=t}^{T-1} (1 + \alpha_i r_s) + \sum_{s=t}^{T-1} (1 - \alpha_i) r_s \quad (1)$$

When  $\alpha_i = 0$ , the individual perceives growth to be linear, fully neglecting compound interest. When  $\alpha_i = 1$ , the person correctly perceives growth to be exponential. Values of  $\alpha_i \in (0, 1)$  generate perceptions between linear and exponential growth. Values  $> 1$  reflect over-estimation of the returns to compounding. To measure exponential-growth bias, we include three hypothetical investment questions in our survey that ask for the value of an asset after a certain amount of

time.<sup>4</sup> For each question  $k$  and each individual  $i$ , we construct a measure of exponential-growth misperception that minimizes the distance between the response and the correct answer informed by equation 1 similarly to Goda et al. (2019). Performance on these questions by OPM employees was similar to the U.S. population: between 29 and 33 percent of survey participants answered the questions within 10% of the correct value as compared to 23 to 31 percent in a representative U.S. sample (Goda et al., 2019).

### 2.4.2 Time preferences

We hypothesize that present-biased individuals are more likely to have gaps between their ideal savings rate and actual savings rate due to procrastination. If so, displaying the gap may be a cue that inspires them to make a change. Though theory does not make a sharp prediction about the direction of change, we explore whether the treatment differentially affects participants based on the degree of their present bias.

We use a “time-staircase” procedure to construct a simple measure of present bias, which we refer to as “Beta”, as well as of the long-run discount factor (“Delta”) in an approach similar to Goda et al. (2019). The method was developed by Falk et al. (2016) for measuring only the long-run discount factor. Staircases have these forms:

**Present-Future Staircase:** Would you rather receive \$100 today or  $\$[X]$  in 12 months?

**Future-Future Staircase:** Would you rather receive \$120 in 12 months or  $\$[Y]$  in 24 months?

Subjects begin with a common value of  $[X]$  or  $[Y]$ . If a subject indicates they prefer the money sooner (later), then the second dollar amount increases (decreases) on the next question.<sup>5</sup> For each staircase, subjects answer five questions, gradually narrowing the interval that contains the indifference point. Since the questions are binary and have parallel structure, they are easily understood and can be answered quickly. Participants were asked these questions for a 12-month (as shown above) and a 6-month time interval, for a total of four sets. We randomize the order of the staircases and use different base values for the different sets of questions (i.e., the Present-Future Staircase always begins with \$100 today and the Future-Future Staircase with \$120 in 12 months)

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<sup>4</sup>An example question is, “An asset has an initial value of \$100 and grows at an interest rate of 10% each period. What is the value of the asset after 20 periods?”

<sup>5</sup>In our survey instrument, the future value  $X$  was always greater than 100 and  $Y$  was always greater than 120.

to minimize the influence of mechanical responses. While this staircase method did not involve real stakes, [Falk et al. \(2016\)](#) show that behavior between a no-stakes and real-stakes version is highly correlated.<sup>6</sup> From these staircases we construct measures of Beta and Delta from the implied indifference point.<sup>7</sup>

### 2.4.3 Financial literacy

Employees with low financial literacy may struggle to make retirement savings decisions, due to not knowing what an appropriate savings rate is. In addition, low financial literacy may create difficulty regarding the process of implementing changes. We hypothesize that employees with low financial literacy would have bigger gaps between their ideal savings rate and their actual savings rate, and that the intervention will have larger treatment effects on those with low financial literacy if the savings tool serves as a substitute for financial capability.

We measure basic financial literacy using the 5-item battery of financial literacy questions developed by [Lusardi and Mitchell \(2011b\)](#) and widely used since then ([Lusardi and Mitchell, 2014](#)). These questions measure understanding of inflation, diversification, compound interest, mortgage payments, and bond prices using multiple choice questions. OPM employees performed well on these questions relative to the U.S. population; percent correct on each of the 5 questions ranged between 39 and 95 percent for OPM employees, compared to 21 to 70 percent for a representative sample of the U.S. population ([Lusardi and Mitchell, 2011b](#)). Similarly, the share of employees who answered all five questions correctly was 30 percent, relative to 10 percent for the U.S. population, suggesting that OPM employees are more financially literate than average. In our subsequent analysis, we use a  $z$ -score of financial literacy standardized within the sample.

### 2.4.4 Factor Analysis

To understand heterogeneity in treatment effects, we take two approaches. First, we look for heterogeneity along theoretically-important dimensions – such as financial literacy, exponential-growth bias (EGB), present bias (beta), educational attainment, prior contribution levels – one at

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<sup>6</sup>The authors find a correlation between the staircase measures and incentivized experimental measures of 0.524. This correlation is close to the test-retest correlation of 0.664 for the incentivized experiment.

<sup>7</sup>We cannot identify the indifference point for those who select the upper bound of the time staircase. In this case, we use the upper bound value plus the difference between that value and the second-to-last value to determine the indifference point. We include a dummy variable for those with these imputed values in the analysis.

a time. Second, we pool information across multiple measures of financial capability and reduce dimensionality by estimating a latent factor and looking at heterogeneous treatment effects between individuals with more or less of this factor.

The first step is to reduce the dimensionality of the heterogeneity by conducting a principal component analysis of the baseline characteristics. Specifically, we include: age in years, gender, years of schooling, race/ethnicity categories, household size, tenure in years, a supervisor status dummy, a permanent tenure status dummy, measured EGB, measured beta, measured delta, and measured financial literacy. We retain six significant factors, and report the rotated factor loading matrix in Table 1.

Table 1: Factor Loading Matrix

Variable	Factor1 Demographics	Factor2 Seniority	Factor3 Financial Capability	Factor4 Time Preference	Factor5 Big Daddy	Factor6 Hispanic Factor	Uniqueness
Age	-0.0753	0.6838	0.0146	0.0648	-0.2091	-0.07	0.4738
Male	0.2269	-0.0046	0.3806	0.046	0.5064	0.0223	0.5446
Years of Schooling	-0.0993	-0.1911	0.7269	-0.0084	-0.1586	0.1145	0.3869
Race = White	0.925	-0.0198	-0.0022	0.0105	-0.0082	-0.2718	0.0699
Race = Hispanic	-0.0756	-0.0451	0.024	0.0178	-0.025	0.9097	0.1632
Race = Black	-0.9478	0.0585	-0.0297	-0.0367	-0.0067	-0.1584	0.071
Household Size	-0.0492	-0.0578	-0.0828	-0.0419	0.8686	-0.0349	0.2299
Tenure(in years)	-0.0802	0.8116	-0.131	0.0262	0.063	-0.0457	0.311
Is Supervisor	0.0577	0.4178	0.3047	-0.0493	0.2453	0.2889	0.5832
Tenure Description = Permanent	-0.0107	0.6444	-0.02	-0.0151	-0.0988	-0.012	0.5741
Std. Alpha	0.0448	0.1002	0.349	-0.0211	0.0972	-0.3106	0.7598
Std. Beta	0.0349	-0.0148	-0.0841	0.8349	-0.074	-0.0388	0.2875
Beta-Delta	0.0313	0.0673	0.1772	0.7921	0.0388	0.0725	0.3289
Financial Literacy	0.1299	0.0207	0.7042	0.1154	0.0648	-0.0656	0.4649
Eigenvalue	2.07686	1.75206	1.50360	1.31937	1.05755	1.04191	

Note: The principal component analysis generated 14 factors, but factors with eigenvalue greater than 1 are retained and reported. Table 1 reports the rotated factor loading matrix from the principal component analysis for the retained factors. Parallel Analysis is performed, as shown in [Appendix D](#)

While these estimated factors are nothing more than a low-dimensional summary of the variation in the data, examining the loadings allows for some meaningful interpretations. For example, the first factor loads primarily on fixed demographic characteristics such as gender and race (and conversely, these dimensions load primarily on this factor). The second retained factor loads primarily on age, length and type of tenure, and supervisory status. We thus interpret this as a composite measure of “Seniority”. We find that the third retained factor loads on years of education, EGB, and financial literacy and interpret this as a composite measure of “Financial Capability”, measuring different aspects of financial sophistication. Finally, the fourth retained factor loads primarily

on the estimated beta and beta $\times$ delta, and so we interpret it as a composite measure of time preference. The remaining factors are less meaningfully interpretable, but still serve to summarize the remaining variation in the data. We use these composite factors to consider heterogeneity in the treatment effect at a higher level of abstraction.

Our pre-analysis plan was registered at the [Social Science Registry AEARCTR-0002129](#). We pre-specified that we would measure the heterogeneous effects of exponential-growth bias, time preferences, and financial literacy but we did not pre-specify the factor analysis or the regressions using the factors. The reader may view these analyses as more exploratory.

### 3 Results

The design of our intervention allows us to investigate three questions. First, we examine whether selection into tool use varies by observable characteristics. Second, we measure the treatment effect among those who clicked the link in the email invitation to use the tool. Finally, we measure how the treatment effect varies with a person’s financial capability to determine whether the treatment is a substitute or complement to financial capability.

#### 3.1 Selection into Tool Use

To examine selection into tool use, we regress tool use on individual characteristics using a logit specification and present our results in Table 2. First, we regress a binary variable that equals 1 for those who use the tool and zero otherwise on mean Alpha, mean Beta, and standardized financial literacy, and show these results in Column 1. None of the coefficients are statistically significant indicating no selection based on the primary variables that we hypothesized would play a role in insufficient retirement saving. We expand the regression to include age, gender, race, education, and household size (Column 2); these coefficients are statistically significant. In Column 3, we layer in employment attributes including total pay, tenure in years, leadership/manager, and tenure status. The only significant effect comes from pre-intervention TSP amount. The effect is highly significant, with an additional standard deviation of TSP amount ( $SD = \$5,707.5$ ) increasing the likelihood of using the tool by  $e^{(5.7075 \times 0.048)} - 1 = 32\%$ . This finding indicates that those who are likely in greatest need of a course correction — those with low saving — are less likely to use the tool.

Table 2: Selection into TOT Sample

	Logit		
	(1)	(2)	(3)
	Tool Participation	Tool Participation	Tool Participation
Tool Participation			
Mean Alpha	0.111 (0.071)	0.107 (0.072)	0.085 (0.073)
Mean Beta	0.393 (0.683)	0.368 (0.699)	0.233 (0.697)
Std. Financial Literacy	0.078 (0.056)	0.044 (0.061)	-0.009 (0.063)
Age		-0.001 (0.006)	-0.009 (0.006)
Male		-0.031 (0.121)	-0.059 (0.125)
White		0.018 (0.292)	0.215 (0.307)
Hispanic		-0.323 (0.390)	-0.171 (0.408)
Black		-0.240 (0.312)	-0.015 (0.325)
Some College or Associate		0.282 (0.198)	0.191 (0.202)
Bachelor		0.240 (0.168)	0.008 (0.177)
Post-Bachelor		0.186 (0.182)	-0.108 (0.202)
Household Size		0.041 (0.045)	0.037 (0.045)
Total Pay			0.003 (0.003)
Tenure in Years			-0.006 (0.009)
Team Leader			0.222 (0.368)
Supervisor or Manager			0.415* (0.247)
Conditional - Tenure Group 2			0.577 (0.494)
Permanent - Tenure Group 1			0.657 (0.454)
Part-Time			0.845 (0.882)
TSP Amount Pre-Rollout (\$1,000/year)			0.048*** (0.013)
Constant	0.252 (0.690)	0.096 (0.849)	-0.575 (1.007)
Mean DV	0.667	0.668	0.668
Observations	1,435	1,393	1,392

Robust standard errors reported. Dependent variable in column heading. The omitted group is female, of other race, with High School education, holding non-supervisory position, and of other tenure group.

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

### 3.2 Treatment on Treated

We next estimate treatment-on-treated (TOT) effects, which represent the differences in contributions among the treatment group and the active control within the subsample of individuals who chose to interact with their version of the tool, rather than the intent-to-treat (ITT) effect among everyone invited to interact with the tool. In most experiments, the econometrician cannot observe which individual in the control group would take up treatment if offered the chance. Our active control design allows us to measure this, creating a particularly strong TOT design and precise estimate. We focus on the TOT effect as it is better powered to detect differences between the conditions.

Using data at the individual-month level, we regress annualized TSP contributions on a post-intervention indicator, the treatment indicator, and the interaction between the two using a difference-in-difference framework. The coefficient on the interaction term is our estimate of the treatment effect for the full treatment relative to the active control. We include year and month fixed effects to control for temporal variation in contributions, and individual fixed effects to control for between-person variation. We cluster standard errors at the person level.

Table 3 shows the main results of the treatment-on-treated analysis. The estimated effect of the treatment is a \$174 increase in contributions per year ( $p=0.021$ ), which represents a 2.5 percent increase in annual contributions compared to the \$7,078 average annual contribution and 0.2 percent of average annual pay (Column 1). We also report the mean of the dependent variable for the estimation sample and a  $p$ -value calculated using permutation inference at the bottom of the table. We do this by randomly re-labeling participants as control and treatment 1,000 times and computing a counterfactual treatment effect for each simulation, which creates a distribution of treatment effects under the null hypothesis. Our estimated treatment effect exceeds all but the top 0.1% of our simulated treatment effects, giving rise to a  $p$ -value smaller than the estimate's  $p$ -value using asymptotic approximation. This effect is similar in magnitude to the effect found in [Goda, Manchester and Sojourner \(2014\)](#), who randomly assigned retirement income projections in the mail to University of Minnesota employees. Their treatment boosted average optional contributions by \$85 annually, which was 3.6 percent of average optional contributions and 0.15 percent of pay.



Table 3: Average Effects and Heterogeneous Effects by Single Dimensions of Heterogeneity (TOT)

	TOT Main		TOT Heterogeneity				
	(1) Overall Sample	(2) Survey Sample	(3) Std. Alpha	(4) Std. Beta	(5) Std. Financial Literacy	(6) TSP Amount per year pre Rollout	(7) Bachelor or Higher
Post $\times$ Full Tool	174.184** (75.621)	120.979 (129.646)	114.466 (129.537)	118.969 (129.367)	132.774 (129.607)	308.069* (174.319)	-210.650 (195.251)
Post $\times$ Attribute			-63.461 (84.566)	120.159 (108.571)	-166.267 (102.292)	0.073*** (0.018)	-179.543 (201.044)
Post $\times$ Full Tool $\times$ Attribute			122.769 (106.152)	-152.713 (131.581)	328.038** (130.793)	-0.022 (0.024)	496.098* (257.274)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	7078.012	7577.489	7577.489	7577.489	7577.489	7577.489	7577.489
Permutation P Value	0.001	0.335					
R-squared	0.089	0.089	0.089	0.089	0.090	0.096	0.090
Observations	151,732	57,744	57,744	57,744	57,744	57,744	57,744

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable is TSP amount. “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (1) reports the estimated TOT effects of all tool users, Col (2) of tool users who also answered the survey. Single-dimension attributes are collected from the survey. Col (3)–(7) report the heterogeneous TOT effects by attributes as specified in the corresponding column heading. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect. For Col (1) and Col (2), p-values from permutation inference of 1,000 times are reported.  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Column 2 reproduces the same specification from Column 1 for the survey-response subsample, which is the sample we use to investigate heterogeneity in the treatment effect. While the TOT estimates in Columns 1 and 2 are similar, the estimate in Column 2 is a bit smaller in magnitude and the standard errors increase with the smaller sample, making the treatment effect no longer statistically significant at conventional levels.

### 3.3 Heterogeneity in Treatment Effects

Next, we estimate heterogeneous treatment effects (Columns 3-7). The coefficient of interest in each column is the three-way interaction (post-intervention  $\times$  treatment group  $\times$  attribute), which may be interpreted as the increase in the treatment effect, relative to the active control, of a one unit increase in the attribute. In Column 3, the attribute is exponential-growth bias; in Column 4, the attribute is the short-run discount rate, Beta; and in Column 5, the attribute is financial literacy. We standardize each of these attributes so a one-unit change corresponds to one standard deviation.

While we find no evidence of heterogeneous treatment effects with respect to exponential-growth bias and present bias, we find evidence of a statistically significant difference in the treatment effect depending on one’s level of financial literacy. The sign of the coefficient indicates that, rather than less financially literate employees benefiting from the increase in information, the treatment has a greater impact on more financially literate individuals, leading them to increase their contributions more. Specifically, one standard deviation higher financial literacy increases the treatment effect by \$328 in annual contributions. Because the treatment leads those with higher levels of financial literacy to make bigger changes in their contributions than those with lower financial literacy, the evidence suggests that the intervention complements rather than substitutes for financial capability.

The lack of significant heterogeneity by exponential-growth bias or present bias ran contrary to our expectation. The intervention was designed to help individuals accurately understand the mapping from retirement account contributions to retirement income. Exponential-growth bias distorts this understanding, tending to lead one to underestimate the future benefits of more-immediate sacrifices [Goda et al. \(2019\)](#). This may be a particularly acute issue for those with naive present bias as well.

In Column 6 we examine heterogeneity in the treatment effect based on pre-intervention con-

tributions. We find no evidence of differences in the treatment effect between those who were contributing different amounts prior to the intervention (Column 6). We estimate heterogeneity by formal educational attainment and find that those with at least a bachelor’s degree exhibit treatment effects that are \$496 greater than those with lower levels of education, though the effect is only marginally significant (Column 7).<sup>8</sup>

Drawing on the latent factors described and estimated in section 2.4.4, in Table 4, we include interaction terms for the three meaningful composite factors — Seniority, Financial Capability, and Time Preference. As before, we include year and month fixed effects to control for temporal variation in contributions and individual fixed effects to control for between-person variation. The coefficients of interest are thus the triple-interactions of  $\text{Post} \times \text{Treatment} \times \text{Factor}$ , which describes how the relative increase in the treatment group over the active control differs for those with a standard deviation higher level of the composite factor. In Columns 1 and 5, the estimated coefficient on the three-way interaction provides evidence that Demographics is not associated with a statistically or economically significant heterogeneity in treatment effects. In Columns 2 and 5 the same can be said about Seniority.<sup>9</sup> In contrast, one standard deviation higher Financial Capability is associated with a \$412 stronger treatment effect. These results are consistent with those in Table 3, where financial literacy and education levels were associated with larger treatment effects. At this greater level of abstraction, we find that more financially capable employees benefit more from the increase in information; that is, the information intervention and financial capability are complements rather than substitutes. Third, we fail to find evidence that Time Preferences mediate the treatment effect. Finally, we include all four factors and their interactions with Treatment and Post simultaneously and find evidence that the only significant interaction is with financial capability and that this significant interaction is evident even when allowing heterogeneity on the other factors.<sup>10</sup>

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<sup>8</sup>Appendix Table A.5 replicates Table 3 but with the outcome in terms of standard deviations of TSP amount. Appendix Table A.8 replicates Table 3 but with the outcome in terms of TSP rate. Intent-to-treat versions of these tables are available in Appendix Table A.4, A.7, and A.10. The main effect is not significant in these analyses but the heterogeneous effects of financial literacy and education are similar in terms of sign and significance, as are the null effects of the other attributes.

<sup>9</sup>To aid interpretation, note that, in the control group, greater Seniority is associated with a \$294 smaller change after the experiment began versus before.

<sup>10</sup>Appendix Table A.6 replicates Table 4 but with the outcome in terms of standard deviations of TSP amount. Appendix Table A.9 replicates Table 4 but with the outcome in terms of TSP rate.

Table 4: Heterogeneous Effects by Factors (TOT)

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)
Post × Full Tool	141.889 (130.840)	75.229 (130.527)	151.798 (131.326)	137.219 (130.473)	173.534 (135.362)	133.807 (131.544)	25.538 (134.771)
Post × Demographics	-105.760 (95.464)						-107.469 (96.001)
Post × Full Tool × Demographics	149.497 (128.685)						157.211 (126.854)
Post × Seniority		-293.914*** (99.988)					-288.275*** (99.769)
Post × Full Tool × Seniority		-38.885 (137.083)					-67.622 (133.333)
Post × Financial Capability			-126.354 (97.740)				-113.895 (96.591)
Post × Full Tool × Financial Capability			411.633*** (132.631)				364.711*** (128.438)
Post × Time Preference				164.910 (109.860)			176.523 (109.173)
Post × Full Tool × Time Preference				-180.815 (133.436)			-180.677 (132.239)
Post × Big Daddy					46.222 (104.020)		57.651 (102.362)
Post × Full Tool × Big Daddy					-101.637 (128.338)		-113.733 (125.478)
Post × Hispanic Factor						-81.289 (93.459)	-78.221 (84.823)
Post × Full Tool × Hispanic Factor						89.919 (108.988)	56.255 (103.873)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	7579.859	7579.859	7579.859	7579.859	7579.859	7579.859	7579.859
F-Statistic	1.350	0.080	9.632	1.836	0.627	0.681	
P-Value	0.246	0.777	0.002	0.176	0.429	0.410	
R-squared	0.089	0.094	0.093	0.092	0.092	0.092	0.107
Observations	56,131	56,131	56,131	56,131	56,131	56,131	56,131

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in column heading. “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Factors are generated from the principal component analysis using single-dimensional attributes from survey as inputs. 6 factors with eigenvalue greater than 1 are retained. Factor loadings are reported in Table 1. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect.\* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

## 4 Discussion

Our results are surprising in several ways. We find that selection into tool use favors those who save more, and who are therefore less likely to need a TSP saving correction. This finding goes against the overall efficacy of the tool as those who are at greatest risk of inadequate retirement savings are those who are least likely to use it. The treatment effect is increasing in financial literacy, education, and our composite Financial Capability metric, generated through exploratory factor analysis. We designed the intervention expecting that behavioral biases likely cause people to make suboptimal retirement-savings decisions, targeting EGB and procrastination, which have been shown in prior work to be associated with lower levels of retirement savings. We therefore hypothesized that an intervention designed to counteract those behavioral biases would improve decisions overall, but more specifically for those who were more biased. However, we found no evidence that either of these biases were correlated with the treatment effect.

Past literature has shown that financial literacy and financial capability is positively correlated with more retirement savings, while controlling for many other variables including income and age (Goda et al., 2019). However, complementarity between the treatment and Financial Capability implies that interventions like the one in this paper may be ineffective at helping employees who are most vulnerable. If this lower savings stems from uninformed decision-making, retirement saving outcomes are likely suboptimal, and so helping individuals who lack financial capability would be a natural policy goal. Our results suggest that simple online retirement-savings tools are not sufficient meet this goal.

We speculate that a certain degree of Financial Capability is necessary to effectively use the online tool. Employees with lower Financial Capability may have been intimidated by the number of steps in the tool, by the financial language (e.g. “TSP”, “catch-up contributions”, “projected income”), or by the self-knowledge required to fill out the entries (e.g. one’s current TSP balance, one’s annual TSP contribution). Past research has shown that financial self-knowledge is low. Bhargava and Conell-Price (2021) find that 20-37% of non-participants in their employer’s 401(k) program mistakenly believed that they had already enrolled. Relatedly, Bhargava et al. (2021) find that cosmetic user-interface design elements can have a large effect on employee savings rates, suggesting that many employees are making decisions in a haphazard or non-deliberative manner.

To help employees with lower Financial Capability, online tools may require better automation whereby the fields in the online tool are auto-populated by the employee’s administrative data. Such integration would lead to fewer steps, less reliance on financial language, and less need for employee self-knowledge. However, it is also possible that more expensive forms of intervention, such as one-on-one sessions or personalized materials, may be necessary to help those with lower Financial Capability.

## 5 Conclusion

We conducted a randomized controlled trial, inviting federal employees to use an online retirement-savings tool. Participants who received projections of their retirement income from their DC plan saved \$174 more annually than those who did not. Selection into the tool favored those who already had higher TSP contributions. The treatment effect was larger for the financially literate and those who were more “Financial Capable,” a factor generated by our factor analysis. This complementarity between the tool and financial capability suggests that similar tools may be effective at helping the well-informed, educated, and financially literate to make retirement-savings decisions, but unlikely to help those who are relatively uninformed, less educated, and less financially literate. Different approaches may be needed to help different populations. One of the strengths of online tools is that they scale well: the marginal cost to the employer or plan manager is near zero. We find evidence of benefits for financially capable workers that may justify those costs. However, these findings suggest that more research and development regarding cost-effective ways to assist those with lower financial capability is needed.

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## (For Online Publication Only)

### Appendix A Additional Results

Table [A.1](#) shows balance of observables by condition assignment for the full sample and for tool users.

Table [A.2](#) shows descriptive statistics for the survey sample.

Table [A.3](#) displays selection into the survey sample. The survey sample is highly selected.

Table [A.4](#) shows both the main intent-to-treat effect, and heterogeneous intent-to-treat effects for Alpha, Beta, financial literacy, TSP amount and education. The main effect is not significant. Only the heterogeneous effect on financial literacy is statistically significant.

Tables [A.5](#), [A.6](#), [A.7](#) replicate Tables [3](#), [4](#), and [A.4](#) except the outcome is in standard deviations of TSP amount.

Tables [A.8](#), [A.9](#) and [A.10](#) replicate Tables [3](#), [4](#), and [A.4](#) except the outcome is in TSP rate.



Table A.1: Descriptive Sstatistics for ITT and TOT Sample

	Assignment				Tool Use				
	(1) All	(2) Partial	(3) Full	(4) Difference	(5) All Tool User	(6) Non-User	(7) Partial User	(8) Full User	(9) Difference
TSP Amount (\$/year)	6274.8 (5721.6)	6287.8 (5783.8)	6262.0 (5660.6)	25.803 (155.366)	7269.9 (6037.8)	5382.0 (5265.6)	7319.5 (6190.1)	7219.2 (5880.0)	100.357 (238.437)
SD Change in TSP Amount	1.107 (1.009)	1.109 (1.020)	1.105 (0.998)	0.005 (0.027)	1.282 (1.065)	0.949 (0.929)	1.291 (1.092)	1.273 (1.037)	0.018 (0.042)
Final TSP Rate	6.899 (5.467)	6.899 (5.611)	6.898 (5.323)	0.000 (0.148)	7.852 (5.869)	6.043 (4.927)	7.870 (6.114)	7.833 (5.610)	0.037 (0.232)
Mean Alpha	0.483 (0.826)	0.472 (0.813)	0.493 (0.838)	-0.021 (0.042)	0.516 (0.836)	0.417 (0.802)	0.480 (0.792)	0.550 (0.875)	-0.069 (0.053)
Mean Beta	1.007 (0.0865)	1.005 (0.0854)	1.008 (0.0875)	-0.003 (0.004)	1.007 (0.0827)	1.006 (0.0935)	1.005 (0.0831)	1.008 (0.0823)	-0.003 (0.005)
Std. Financial Literacy	-0.0753 (1.019)	-0.0844 (1.023)	-0.0664 (1.015)	-0.018 (0.053)	-0.0445 (0.995)	-0.138 (1.065)	-0.0400 (1.008)	-0.0487 (0.984)	0.009 (0.064)
Total Pay (in Thousand)	85.99 (31.62)	86.08 (31.74)	85.90 (31.50)	0.180 (0.859)	88.61 (31.77)	83.64 (31.30)	88.71 (32.48)	88.51 (31.04)	0.195 (1.255)
Age	45.73 (10.70)	45.80 (10.69)	45.65 (10.70)	0.144 (0.290)	46.72 (10.43)	44.83 (10.86)	46.75 (10.53)	46.69 (10.33)	0.058 (0.412)
Gender	0.429 (0.495)	0.428 (0.495)	0.429 (0.495)	-0.001 (0.013)	0.443 (0.497)	0.416 (0.493)	0.444 (0.497)	0.441 (0.497)	0.003 (0.020)
Bachelor or Higher	0.654 (0.476)	0.659 (0.474)	0.649 (0.477)	0.010 (0.013)	0.658 (0.475)	0.651 (0.477)	0.679 (0.467)	0.636 (0.481)	0.043* (0.019)
White	0.658 (0.474)	0.653 (0.476)	0.664 (0.473)	-0.011 (0.013)	0.684 (0.465)	0.635 (0.481)	0.688 (0.464)	0.680 (0.467)	0.008 (0.018)
Observations	5,426	2,696	2,730	5,426	2,566	2,860	1,297	1,269	2,566
Chi-Sqaured				2.42					2.49
P-Value				0.97					0.96

Note: Summary statistics of all outcome variables one month before the intervention and single-dimension attributes are reported. Single-dimension attributes are obtained from the survey. Standard deviation in parentheses below. Selected sample in column heading “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (5) reports the difference between active control and treatment group in the ITT sample, with join significant test statistics reported at the bottom. Col (9) reports the difference between active control and treatment group in the TOT sample. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table A.2: Descriptive Statistics by Survey Participation

	(1) All	(2) Survey Non-Completers	(3) Survey Completer	(4) Difference
TSP Amount (\$/year)	6274.0 (5724.1)	5939.1 (5537.6)	7205.4 (6119.9)	-1266.219*** (175.365)
SD Change in TSP Amount	1.107 (1.010)	1.048 (0.977)	1.271 (1.080)	-0.223*** (0.031)
Final TSP Rate	6.895 (5.465)	6.568 (5.268)	7.801 (5.885)	-1.233*** (0.167)
Total Pay (in Thousand)	85.99 (31.62)	85.30 (31.60)	87.90 (31.60)	-2.598** (0.973)
Age	45.73 (10.70)	45.18 (10.65)	47.24 (10.69)	-2.052*** (0.328)
Gender	0.429 (0.495)	0.424 (0.494)	0.442 (0.497)	-0.018 (0.015)
Bachelor or Higher	0.654 (0.476)	0.651 (0.477)	0.663 (0.473)	-0.013 (0.015)
White	0.658 (0.474)	0.642 (0.479)	0.704 (0.457)	-0.062*** (0.015)
Observations	5,426	3,991	1,435	5,426
Chi-Squared				62.39
P-Value				0.00

Note: Summary statistics for all outcome variables one month before the intervention and demographics are reported by survey participation. Robust standard errors reported. Sample selection in column heading. Col (5) reports the difference between survey participants and non-participants, with joint significant test statistics reported at the bottom. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .



Table A.3: Selection into Survey Sample

	Logit	
	(1) In Survey Sample	(2) In Survey Sample
In Survey Sample		
Age	-0.003*** (0.001)	0.001 (0.001)
Male	0.355*** (0.017)	0.356*** (0.017)
White	0.351*** (0.037)	0.359*** (0.037)
Hispanic	-0.106** (0.048)	-0.077 (0.049)
Black	0.202*** (0.039)	0.254*** (0.040)
Some College or Associate	0.503*** (0.028)	0.492*** (0.029)
Bachelor	0.105*** (0.021)	0.103*** (0.023)
Post-Bachelor	0.315*** (0.024)	0.300*** (0.027)
Household Size	0.054*** (0.006)	0.061*** (0.007)
Total Pay		-0.002*** (0.000)
Tenure in Years		-0.019*** (0.001)
Team Leader		0.133*** (0.047)
Supervisor or Manager		-0.001 (0.031)
Conditional - Tenure Group 2		-0.459*** (0.069)
Permanent - Tenure Group 1		-0.104* (0.063)
Part-Time		1.421*** (0.186)
Full-Time		1.572*** (0.169)
Constant	0.807*** (0.059)	-0.490*** (0.188)
Mean DV	0.806	0.806
Observations	103,607	103,607

Note: Robust standard errors reported. Dependent variable in column heading. The omitted group is female, of other race, with High School education, holding non-supervisory position, and of other tenure group.

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table A.4: Effect of the Treatment (ITT) on TSP Amount

	ITT Main		ITT Heterogeneity				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Overall Sample	Survey Sample	Std. Alpha	Std. Beta	Std. Financial Literacy	TSP Amount per year pre Rollout	Bachelor or Higher
Post $\times$ Full Tool	61.055 (48.990)	134.103 (100.994)	131.192 (100.774)	134.080 (100.901)	151.680 (101.817)	285.584** (135.674)	-89.439 (148.638)
Post $\times$ Attribute			41.775 (74.787)	30.028 (73.575)	-125.891* (75.388)	0.081*** (0.014)	
Post $\times$ Full Tool $\times$ Attribute			80.896 (92.855)	21.494 (92.759)	238.383** (99.264)	-0.021 (0.020)	
Post $\times$ Attribute=1							-90.545 (147.613)
Post $\times$ Attribute=1 $\times$ Full Tool							337.035* (198.862)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	6188.494	7016.741	7016.741	7016.741	7016.741	7016.741	7016.741
F-Statistic			0.759	0.054	5.767	1.089	2.872
P-Value			0.384	0.817	0.016	0.297	0.090
R-squared	0.069	0.072	0.073	0.072	0.073	0.081	0.073
Observations	318,873	85,974	85,974	85,974	85,974	85,974	85,974

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in title: “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (1) reports the estimated ITT effects, Col (2) of who answered the survey. Single-dimension attributes are collected from the survey. Col (3)–(7) reports the heterogeneous ITT effects by attributes as specified in the corresponding column heading. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect.  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table A.5: Effect of the Treatment (TOT) on SD Change in TSP Amount

	TOT Main		TOT Heterogeneity				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Overall Sample	Survey Sample	Std. Alpha	Std. Beta	Std. Financial Literacy	TSP Amount per year pre Rollout	Bachelor or Higher
Post $\times$ Full Tool	0.031** (0.013)	0.021 (0.023)	0.020 (0.023)	0.021 (0.023)	0.023 (0.023)	0.054* (0.031)	-0.037 (0.034)
Post $\times$ Attribute			-0.011 (0.015)	0.021 (0.019)	-0.029 (0.018)	0.000*** (0.000)	-0.032 (0.035)
Post $\times$ Full Tool $\times$ Attribute			0.022 (0.019)	-0.027 (0.023)	0.058** (0.023)	-0.000 (0.000)	0.088* (0.045)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	1.248533	1.336639	1.336639	1.336639	1.336639	1.336639	1.336639
Permutation P Value	0.000	0.348					
R-squared	0.089	0.089	0.089	0.089	0.090	0.096	0.090
Observations	151,732	57,744	57,744	57,744	57,744	57,744	57,744

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in title. “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (1) reports the estimated TOT effects of tool users, Col (2) of tool users who also answered the survey. Single-dimension attributes are collected from the survey. Col (3)–(7) reports the heterogeneous TOT effects by attributes as specified in the corresponding column heading. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect. For Col (1) and Col (2), p-values from permutation inference of 1,000 times are reported. p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

Table A.6: Heterogeneous Effects by Factors (TOT) on SD Change in TSP Amount

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount
Post × Full Tool	0.025 (0.023)	0.013 (0.023)	0.027 (0.023)	0.024 (0.023)	0.031 (0.024)	0.024 (0.023)	0.005 (0.024)
Post × Demographics	-0.019 (0.017)						-0.019 (0.017)
Post × Full Tool × Demographics	0.026 (0.023)						0.028 (0.022)
Post × Seniority		-0.052*** (0.018)					-0.051*** (0.018)
Post × Full Tool × Seniority		-0.007 (0.024)					-0.012 (0.024)
Post × Financial Capability			-0.022 (0.017)				-0.020 (0.017)
Post × Full Tool × Financial Capability			0.073*** (0.023)				0.064*** (0.023)
Post × Time Preference				0.029 (0.019)			0.031 (0.019)
Post × Full Tool × Time Preference				-0.032 (0.024)			-0.032 (0.023)
Post × Big Daddy					0.008 (0.018)		0.010 (0.018)
Post × Full Tool × Big Daddy					-0.018 (0.023)		-0.020 (0.022)
Post × Hispanic Factor						-0.014 (0.016)	-0.014 (0.015)
Post × Full Tool × Hispanic Factor						0.016 (0.019)	0.010 (0.018)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	1.337	1.337	1.337	1.337	1.337	1.337	1.337
F-Statistic	1.350	0.080	9.632	1.836	0.627	0.681	
P-Value	0.246	0.777	0.002	0.176	0.429	0.410	
R-squared	0.089	0.094	0.093	0.092	0.092	0.092	0.107
Observations	56,131	56,131	56,131	56,131	56,131	56,131	56,131

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in column heading. “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Factors are generated from the principal component analysis using single-dimensional attributed obtained from survey. 6 factors with eigenvalue greater than 1 are retained. Factor loadings are reported in Table 1. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect.\* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

Table A.7: Effect of the Treatment (ITT) on SD Change in TSP Amount

	ITT Main		ITT Heterogeneity				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Overall Sample	Survey Sample	Std. Alpha	Std. Beta	Std. Financial Literacy	TSP Amount per year pre Rollout	Bachelor or Higher
Post $\times$ Full Tool	0.011 (0.009)	0.024 (0.018)	0.023 (0.018)	0.024 (0.018)	0.027 (0.018)	0.050** (0.024)	-0.016 (0.026)
Post $\times$ Attribute			0.007 (0.013)	0.005 (0.013)	-0.022* (0.013)	0.000*** (0.000)	
Post $\times$ Full Tool $\times$ Attribute			0.014 (0.016)	0.004 (0.016)	0.042** (0.018)	-0.000 (0.000)	
Post $\times$ Attribute=1							-0.016 (0.026)
Post $\times$ Attribute=1 $\times$ Full Tool							0.059* (0.035)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	1.092	1.238	1.238	1.238	1.238	1.238	1.238
F-Statistic			0.759	0.054	5.767	1.089	2.872
P-Value			0.384	0.817	0.016	0.297	0.090
R-squared	0.069	0.072	0.073	0.072	0.073	0.081	0.073
Observations	318,873	85,974	85,974	85,974	85,974	85,974	85,974

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in title: “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (1) reports the estimated ITT effects, Col (2) of who also answered the survey. Single-dimension attributes are collected from the survey. Col (3)–(7) reports the heterogeneous ITT effects by attributes as specified in the corresponding column heading. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect.  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .



Table A.8: Effect of the Treatment (TOT) on TSP Rate

	TOT Main		TOT Heterogeneity				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Overall Sample	Survey Sample	Std. Alpha	Std. Beta	Std. Financial Literacy	TSP Amount per year pre Rollout	Bachelor or Higher
Post $\times$ Full Tool	0.145 (0.088)	0.119 (0.162)	0.112 (0.163)	0.116 (0.163)	0.130 (0.162)	0.453* (0.233)	-0.372 (0.289)
Post $\times$ Attribute			-0.061 (0.106)	0.130 (0.157)	-0.325** (0.136)	0.000** (0.000)	-0.667** (0.291)
Post $\times$ Full Tool $\times$ Attribute			0.125 (0.128)	-0.175 (0.175)	0.412** (0.171)	-0.000 (0.000)	0.727** (0.349)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	7.687612	8.166443	8.166443	8.166443	8.166443	8.166443	8.166443
Permutation P Value	0.051	0.452					
R-squared	0.023	0.024	0.024	0.024	0.025	0.026	0.025
Observations	151,732	57,744	57,744	57,744	57,744	57,744	57,744

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in title. “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (1) reports the estimated TOT effects of tool users, Col (2) of tool users who also answered the survey. Single-dimension attributes are collected from the survey. Col (3)–(7) reports the heterogeneous TOT effects by attributes as specified in the corresponding column heading. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect. For Col (1) and Col (2), p-values from permutation inference of 1,000 times are reported. p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

Table A.9: Heterogeneous Effects by Factors (TOT) on TSP Rate

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Final TSP Rate	Final TSP Rate	Final TSP Rate	Final TSP Rate	Final TSP Rate	Final TSP Rate	Final TSP Rate
Post × Full Tool	0.148 (0.164)	0.010 (0.167)	0.136 (0.167)	0.133 (0.164)	0.166 (0.166)	0.145 (0.165)	-0.070 (0.181)
Post × Demographics	-0.075 (0.102)						-0.079 (0.100)
Post × Full Tool × Demographics	0.147 (0.142)						0.163 (0.141)
Post × Seniority		-0.456*** (0.149)					-0.428*** (0.146)
Post × Full Tool × Seniority		0.078 (0.190)					0.025 (0.186)
Post × Financial Capability			-0.375** (0.148)				-0.357** (0.145)
Post × Full Tool × Financial Capability			0.517*** (0.187)				0.465** (0.180)
Post × Time Preference				0.178 (0.151)			0.203 (0.151)
Post × Full Tool × Time Preference				-0.183 (0.171)			-0.202 (0.172)
Post × Big Daddy					0.153 (0.119)		0.152 (0.114)
Post × Full Tool × Big Daddy					-0.200 (0.147)		-0.190 (0.142)
Post × Hispanic Factor						-0.097 (0.096)	-0.083 (0.084)
Post × Full Tool × Hispanic Factor						0.070 (0.118)	0.031 (0.111)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	8.176	8.176	8.176	8.176	8.176	8.176	8.176
F-Statistic	1.078	0.169	7.665	1.141	1.845	0.349	
P-Value	0.299	0.682	0.006	0.286	0.175	0.555	
R-squared	0.024	0.029	0.027	0.025	0.025	0.025	0.038
Observations	56,131	56,131	56,131	56,131	56,131	56,131	56,131

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in column heading. “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Factors are generated from the principal component analysis using single-dimensional attributes from survey as inputs. 6 factors with eigenvalue greater than 1 are retained. Factor loadings are reported in Table 1. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect.\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table A.10: Effect of the Treatment (ITT) on TSP Rate

	ITT Main		ITT Heterogeneity				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Overall Sample	Survey Sample	Std. Alpha	Std. Beta	Std. Financial Literacy	TSP Amount per year pre Rollout	Bachelor or Higher
Post $\times$ Full Tool	0.033 (0.055)	0.103 (0.122)	0.101 (0.122)	0.103 (0.123)	0.126 (0.122)	0.402** (0.173)	-0.238 (0.206)
Post $\times$ Attribute			0.051 (0.089)	0.037 (0.104)	-0.266*** (0.098)	0.000*** (0.000)	
Post $\times$ Full Tool $\times$ Attribute			0.073 (0.108)	0.018 (0.120)	0.319*** (0.123)	-0.000 (0.000)	
Post $\times$ Attribute=1							-0.499** (0.203)
Post $\times$ Attribute=1 $\times$ Full Tool							0.515** (0.256)
Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mean DV	6.848	7.707	7.707	7.707	7.707	7.707	7.707
F-Statistic			0.454	0.023	6.723	2.399	4.055
P-Value			0.501	0.879	0.010	0.122	0.044
R-squared	0.014	0.016	0.016	0.016	0.017	0.019	0.017
Observations	318,873	85,974	85,974	85,974	85,974	85,974	85,974

Note: Robust standard errors in parenthesis and clustered at person level. Dependent variable in title: “Partial” refers to the tool in the active control condition and “Full” refers to the tool in the treatment condition. Col (1) reports the estimated ITT effects, Col (2) of who also answered the survey. Single-dimension attributes are collected from the survey. Col (3)–(7) reports the heterogeneous ITT effects by attributes as specified in the corresponding column heading. All specifications include post dummy, year fixed effect, month fixed effect, and individual fixed effect.  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

## Appendix B TOT Effects by Assumptions

Table B.11: Heterogeneous Effects by Assumptions (TOT) on TSP Amount

	(1)	(2)	(3)	(4)	(5)
	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)	TSP Amount (\$/year)
Post × LR-HL Full Tool	287.964** (131.179)				
Post × HR-HL Full Tool	3.149 (104.879)				
Post × LR-LL Full Tool	211.459* (118.889)				
Post × HR-LL Full Tool	211.512 (129.502)				
Post × LR-HL Partial Tool		50.926 (105.181)			
Post × LR-HL Full Tool		314.025** (142.692)			
Post × HR-HL Full Tool		29.210 (118.974)			
Post × LR-LL Full Tool		237.520* (131.488)			
Post × HR-LL Full Tool		237.573* (141.156)			
Post × Full Tool			248.594*** (95.801)	211.489** (95.195)	280.937*** (107.046)
Post × Full Tool × High Return			-147.862 (108.815)		-144.777 (109.623)
Post × Full Tool × High Lifestyle				-73.336 (108.891)	-66.632 (109.658)
Year F.E.	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes
Omitted	All Partial	LR-LL Partial	All Partial	LL Partial	LR-LL Partial
Assumptions Type	Separating	Separating	Pooling	Pooling	Pooling
Mean DV	7078.012	7078.012	7078.012	7078.012	7078.012
R-squared	0.090	0.090	0.089	0.089	0.090
Observations	151,732	151,732	151,732	151,732	151,732

Notes: The active control group (“Partial Tool”) were assigned two assumptions: Low Return, Low Lifestyle (LR-LL) and Low Return, High Lifestyle (LR-HL). The treatment group (“Full Tool”) were assigned four assumptions: Low Return-Low Lifestyle (LR-LL), Low Return-High Lifestyle (LR-HL), High Return-Low Lifestyle (HR-LL), and High Return-High Lifestyle (HR-HL) . In Col (3)–(5), assumptions are pooled by return and by lifestyle. Robust standard errors in parentheses and clustered on ID. Dependent variables as indicated in column heading. All specifications also include controls for post dummy, year fixed effect, month fixed effect, and individual fixed effect. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table B.12: Heterogeneous Effects by Assumptions (TOT) on SD change in TSP Amount

	(1)	(2)	(3)	(4)	(5)
	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount	SD Change in TSP Amount
Post × LR-HL Full Tool	0.051** (0.023)				
Post × HR-HL Full Tool	0.001 (0.019)				
Post × LR-LL Full Tool	0.037* (0.021)				
Post × HR-LL Full Tool	0.037 (0.023)				
Post × LR-HL Partial Tool		0.009 (0.019)			
Post × LR-HL Full Tool		0.055** (0.025)			
Post × HR-HL Full Tool		0.005 (0.021)			
Post × LR-LL Full Tool		0.042* (0.023)			
Post × HR-LL Full Tool		0.042* (0.025)			
Post × Full Tool			0.044*** (0.017)	0.037** (0.017)	0.050*** (0.019)
Post × Full Tool × High Return			-0.026 (0.019)		-0.026 (0.019)
Post × Full Tool × High Lifestyle				-0.013 (0.019)	-0.012 (0.019)
Year F.E.	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes
Omitted	All Partial	LR-LL Partial	All Partial	LL Partial	LR-LL Partial
Assumptions Type	Separating	Separating	Pooling	Pooling	Pooling
Mean DV	1.249	1.249	1.249	1.249	1.249
R-squared	0.090	0.090	0.089	0.089	0.090
Observations	151,732	151,732	151,732	151,732	151,732

Notes: The active control group (“Partial Tool”) were assigned two assumptions: Low Return, High Lifestyle (LR-HL) and Low Return, Low Lifestyle (LR-LL). The treatment group (“Full Tool”) were assigned four assumptions: Low Return-High Lifestyle (LR-HL), Low Return-Low Lifestyle (LR-LL), High Return-High Lifestyle (HR-HL), and High Return-Low Lifestyle (HR-LL). In Col (3)-(5), assumptions are pooled by return and by lifestyle. Robust standard errors in parentheses and clustered on ID. Dependent variables as indicated in column heading. All specifications also include controls for post dummy, year fixed effect, month fixed effect, and individual fixed effect. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table B.13: Heterogeneous Effects by Assumptions (TOT) on TSP Rate

	(1)	(2)	(3)	(4)	(5)
	Final TSP Rate	Final TSP Rate	Final TSP Rate	Final TSP Rate	Final TSP Rate
Post × LR-HL Full Tool	0.300* (0.159)				
Post × HR-HL Full Tool	-0.060 (0.119)				
Post × LR-LL Full Tool	0.218* (0.128)				
Post × HR-LL Full Tool	0.139 (0.139)				
Post × LR-HL Partial Tool		0.010 (0.131)			
Post × LR-HL Full Tool		0.305* (0.172)			
Post × HR-HL Full Tool		-0.055 (0.136)			
Post × LR-LL Full Tool		0.223 (0.144)			
Post × HR-LL Full Tool		0.144 (0.154)			
Post × Full Tool			0.258** (0.112)	0.180* (0.105)	0.286** (0.118)
Post × Full Tool × High Return			-0.225* (0.119)		-0.222* (0.121)
Post × Full Tool × High Lifestyle				-0.070 (0.119)	-0.059 (0.120)
Year F.E.	Yes	Yes	Yes	Yes	Yes
Month F.E.	Yes	Yes	Yes	Yes	Yes
Individual F.E.	Yes	Yes	Yes	Yes	Yes
Omitted	All Partial	LR-LL Partial	All Partial	LL Partial	LR-LL Partial
Assumptions Type	Separating	Separating	Pooling	Pooling	Pooling
Mean DV	7.688	7.688	7.688	7.688	7.688
R-squared	0.024	0.024	0.024	0.024	0.024
Observations	151,732	151,732	151,732	151,732	151,732

Notes: The active control group (“Partial Tool”) were assigned two assumptions: Low Return, Low Lifestyle (LR-LL) and Low Return, High Lifestyle (LR-HL). The treatment group (“Full Tool”) were assigned four assumptions: Low Return-Low Lifestyle (LR-LL), Low Return-High Lifestyle (LR-HL), High Return-Low Lifestyle (HR-LL), and High Return-High Lifestyle (HR-HL) . In Col (3)–(5), assumptions are pooled by return and by lifestyle. Robust standard errors in parentheses and clustered on ID. Dependent variables as indicated in column heading. All specifications also include controls for post dummy, year fixed effect, month fixed effect, and individual fixed effect. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

## Appendix C Sample Schematic

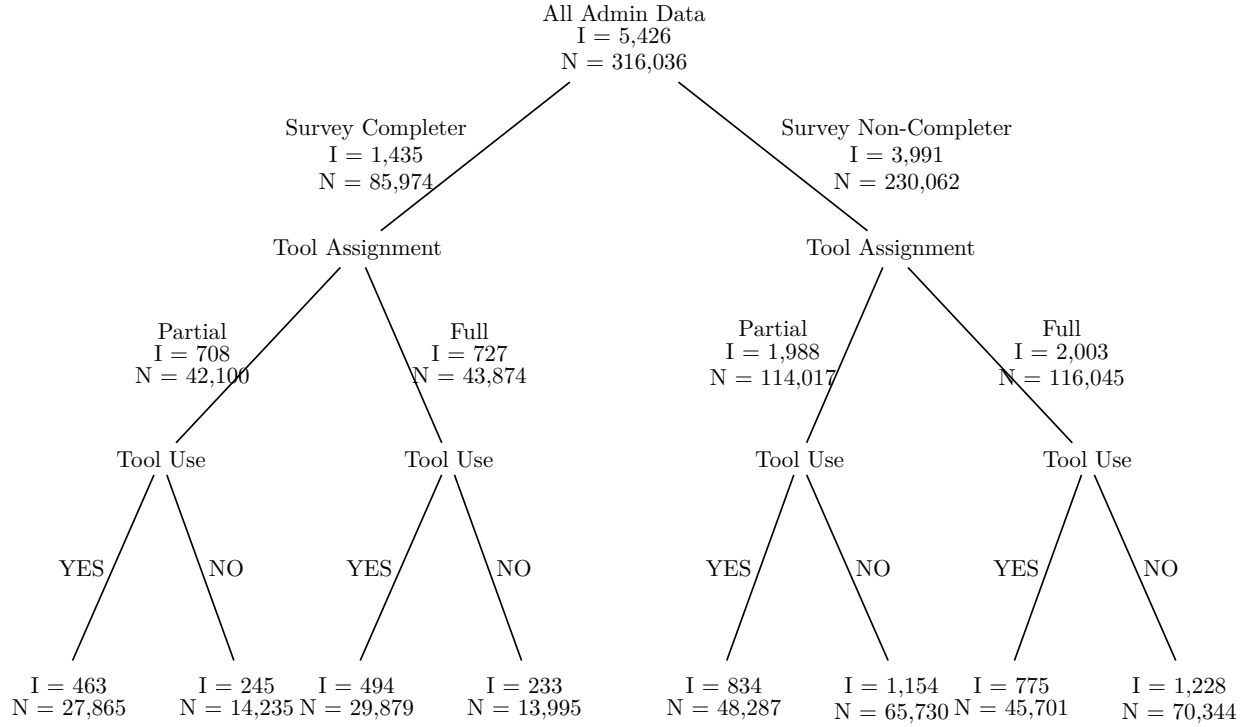


Figure C.1: Sample Schematics

Note:

$I$  - the number of unique individuals in the corresponding node

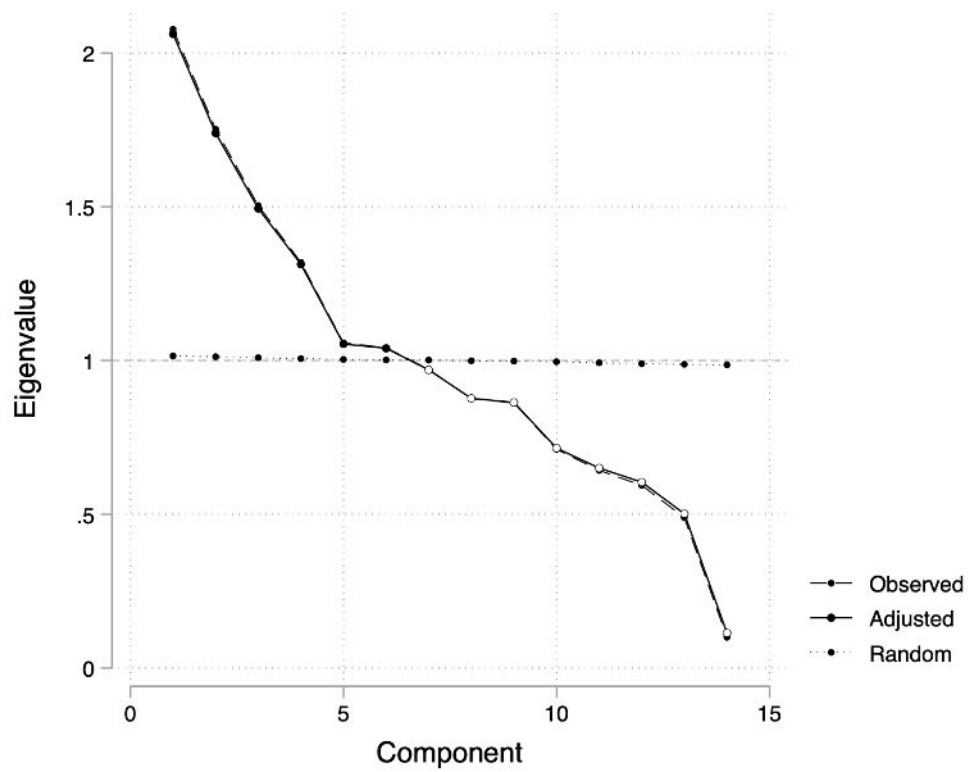
$N$  - the number of observations, the unit of observation is bimonthly paychecks for each individual.

Survey Non-completers include individuals who did not answer all five questions as well as individuals who did not participate in the survey at all.




## Appendix D Parallel Analysis

Figure D.2: Parallel Analysis for Factors



## Appendix E Screenshots

Figure E.3: Step 1

 **Ballpark Savings Estimate**  
Are you saving enough for retirement?

[More Info](#) [Reset Data and Start Again](#)

Step 1 ▶ Step 2 ▶ Step 3 ▶ Step 4 ▶ Step 5 ▶ Step 6 ▶ View Your Plan ▶ Adjust Your Plan ▶ What to Do Next

**Let's get started**

What is your date of birth? Month: -- Year: --


When did you start working for the Federal government? (Service Computation Date) Month: -- Year: --

Current Annual Salary \$

Expected Retirement Age

[Report Issue](#) ◀ ▶

Figure E.4: Step 2



## Ballpark Savings Estimate

Are you saving enough for retirement?

[More info](#)[Reset Data and Start Again](#)

Step 1

Step 2

Step 3

Step 4

Step 5

Step 6


View Your Plan

Adjust Your Plan

What to Do Next


### What lifestyle would you like in retirement?

Select your desired lifestyle in retirement. This will set your retirement income goal.




**70% of Income**

Expect to have **lower spending** in retirement as compared to today.




**85% of Income**

Expect to have **similar spending** in retirement as compared to today.



**100% of Income**

Expect to have **higher spending** in retirement as compared to today.



**115% of Income**

Or enter other amount  (%)

[Report Issue](#)





Figure E.5: Step 3

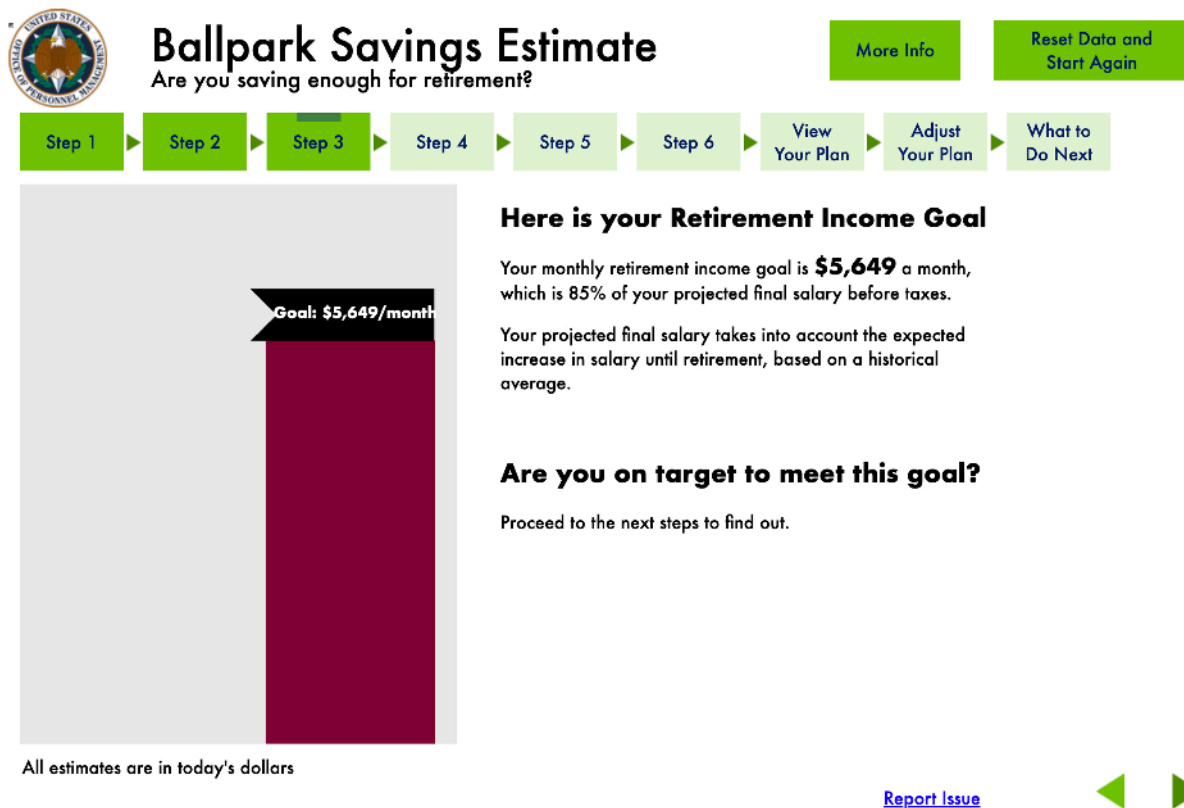
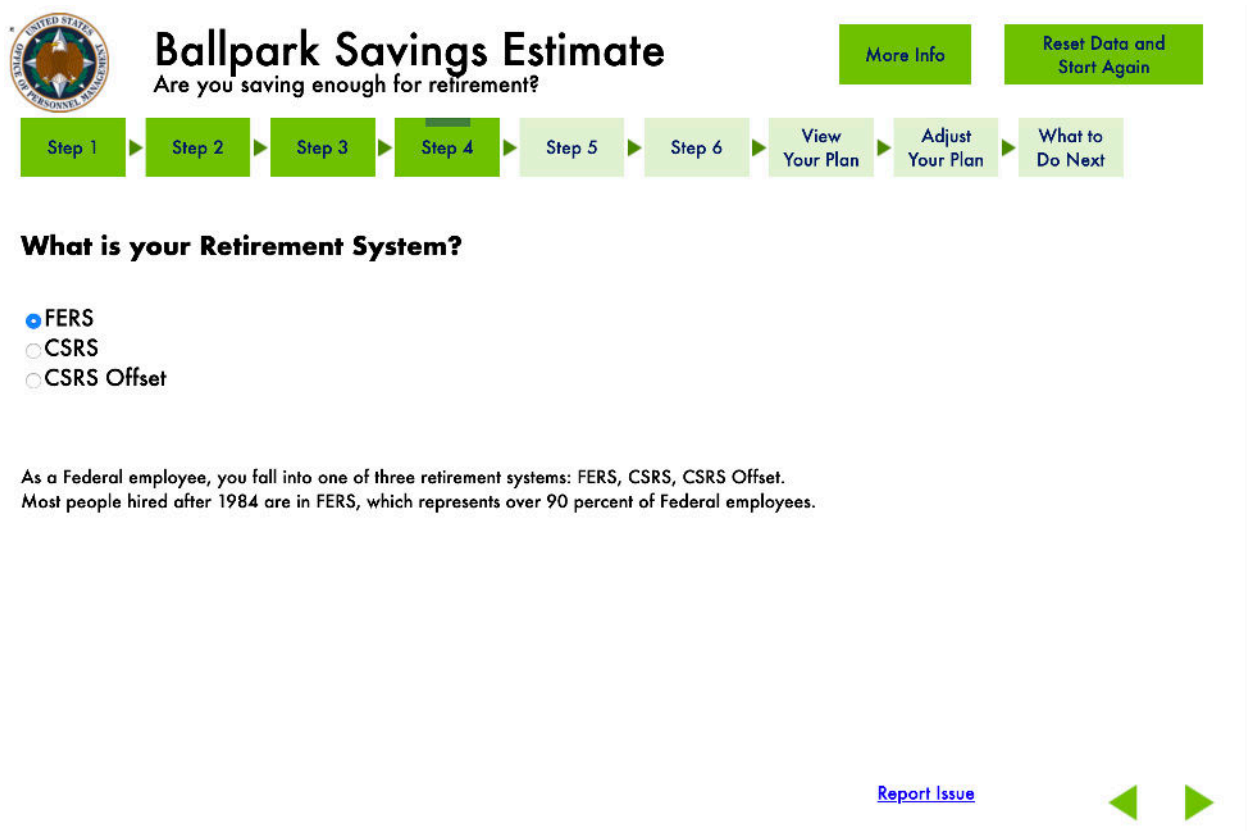



Figure E.6: Step 4



 **Ballpark Savings Estimate**  
Are you saving enough for retirement?

[More Info](#) [Reset Data and Start Again](#)

Step 1 ▶ Step 2 ▶ Step 3 ▶ **Step 4** ▶ Step 5 ▶ Step 6 ▶ View Your Plan ▶ Adjust Your Plan ▶ What to Do Next


**What is your Retirement System?**

☒ FERS  
☐ CSRS  
☐ CSRS Offset

As a Federal employee, you fall into one of three retirement systems: FERS, CSRS, CSRS Offset.  
Most people hired after 1984 are in FERS, which represents over 90 percent of Federal employees.

[Report Issue](#) ◀ ▶

Figure E.7: Step 5



## Ballpark Savings Estimate

Are you saving enough for retirement?

[More Info](#)[Reset Data and Start Again](#)

Step 1

Step 2

Step 3

Step 4

Step 5

Step 6

View Your Plan

Adjust Your Plan

What to Do Next

### What are your current retirement savings?

Federal employees can save additional income for retirement through the Thrift Savings Program (TSP).

Enter Current TSP Account Balance \$

Enter Your TSP Contribution ☒ Percent  %  
☐ Dollar \$   
Max: \$18,500/year or \$712/pay period

#### Annual TSP Catch-up Contribution


Max: 6000

Min: 0

Enter Additional Retirement Savings Balance \$

[Report Issue](#)

Figure E.8: Step 6



## Ballpark Savings Estimate

Are you saving enough for retirement?

[More Info](#)[Reset Data and Start Again](#)

Step 1

Step 2

Step 3

Step 4

Step 5

Step 6

View Your Plan

Adjust Your Plan

What to Do Next

**Do you have other sources of retirement income?**

☒ I expect to receive Social Security benefits.

Enter Expected Monthly Social Security Benefits \$

Need help? [Estimate my Social Security benefit](#)

☐ I expect to work after retirement.

☐ I expect an additional pension.

[Report Issue](#)





Figure E.9: Step 7

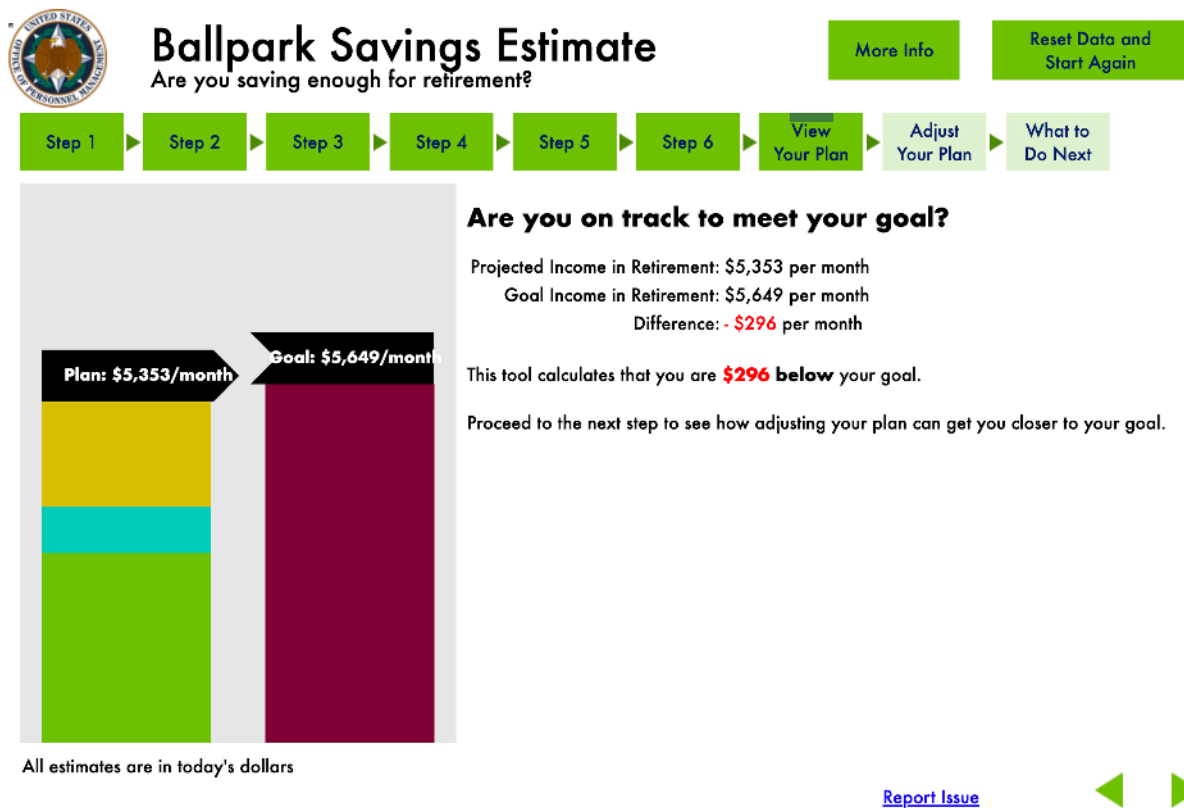




Figure E.10: Step 8

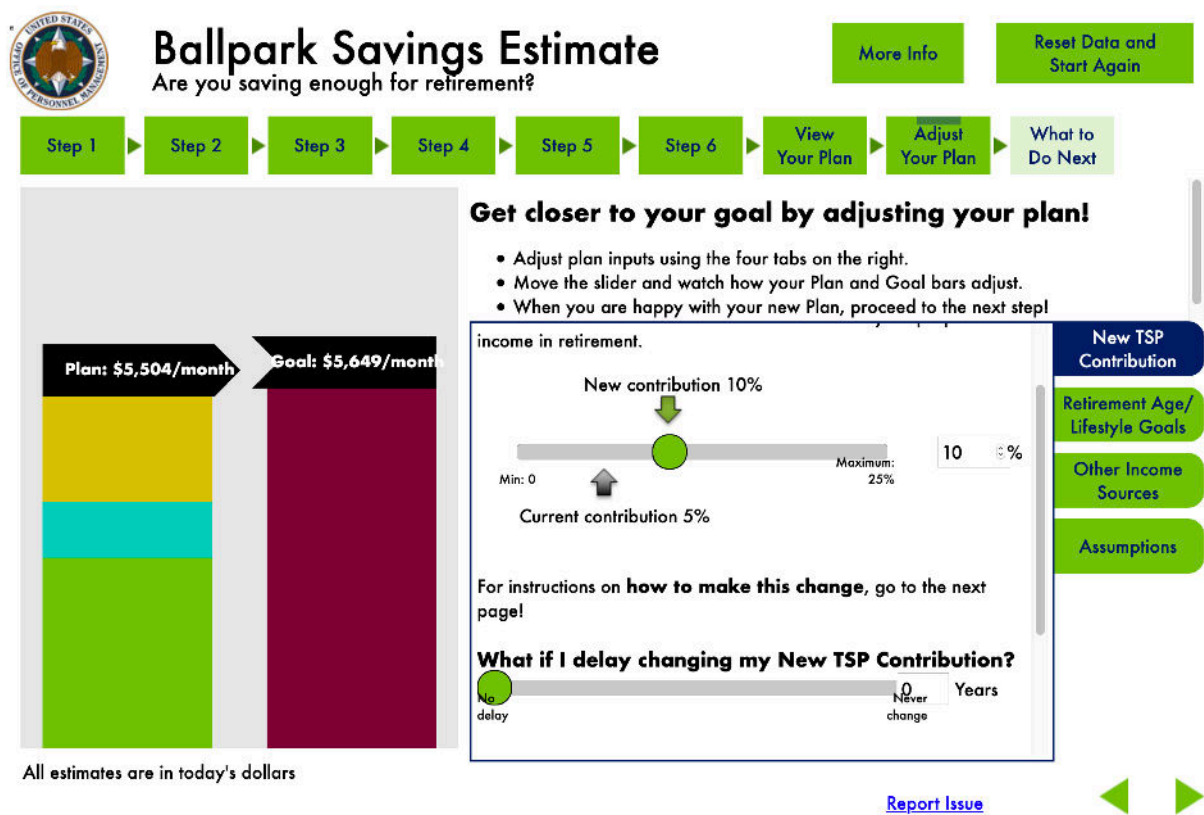


Figure E.11: Step 9

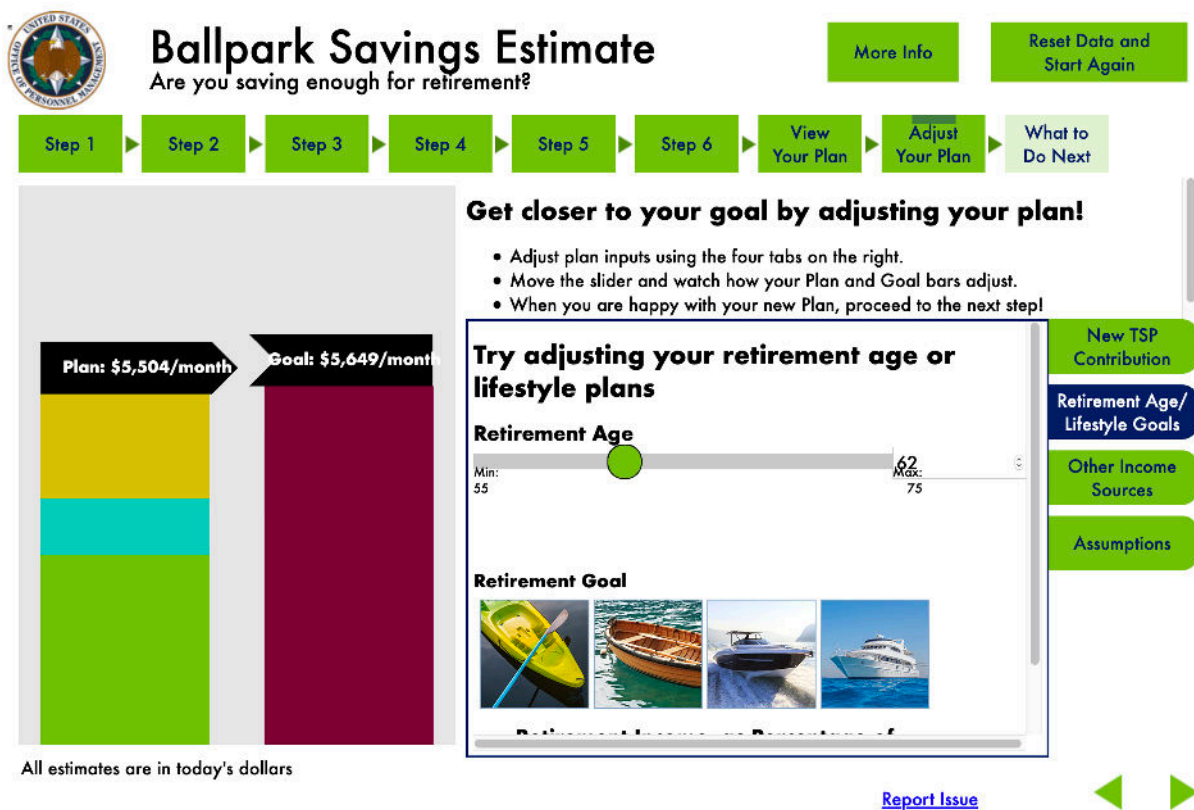


Figure E.12: Step 10



Figure E.13: Step 11

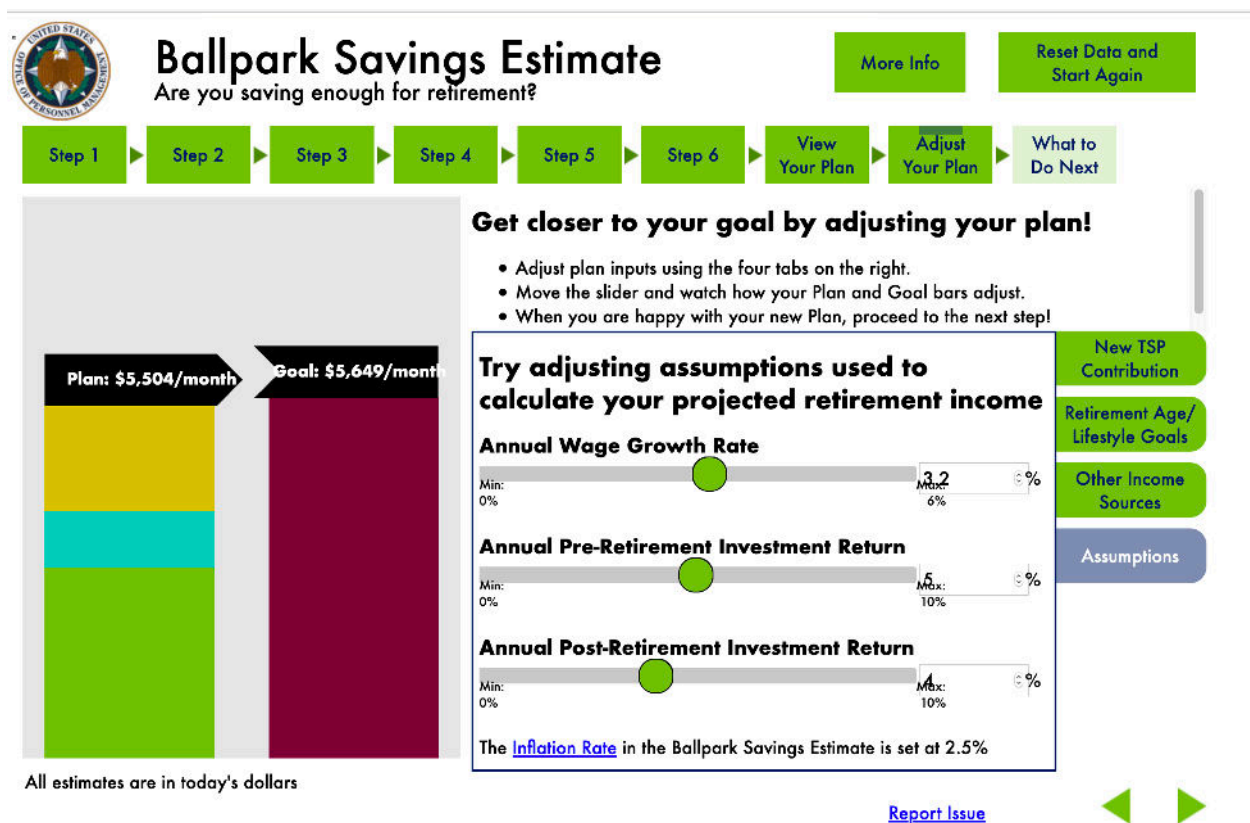



Figure E.14: Step 12



## Ballpark Savings Estimate

Are you saving enough for retirement?

[More Info](#)
[Reset Data and Start Again](#)

Step 1

Step 2

Step 3

Step 4

Step 5

Step 6

View Your Plan

Adjust Your Plan

What to Do Next

**Here is a summary of your Current Saving Plan and your proposed New Saving Plan based on using this tool:**

**Current Saving Plan**

Your Current TSP Contribution:	5% per pay period
Projected Income in Retirement:	\$5,353 per month
Goal Income in Retirement:	\$5,649 per month
Difference:	<span style="color: red;">-\$296</span> per month

This tool calculates that you are \$296 below your goal with your Current Saving Plan.

➤

**New Saving Plan**

Your New TSP Contribution:	10% per pay period
Projected Income in Retirement:	\$5,504 per month
Goal Income in Retirement:	\$5,649 per month
Difference:	<span style="color: red;">-\$145</span> per month

This tool calculates that you are \$145 below your goal with your New Saving Plan.

Print this plan to keep for your records [Print](#)

### Change your TSP contribution now! Here's how:

Sign into your agency's electronic payroll system and select the "Thrift Savings Plan" option. You can contribute a percentage of your salary or a fixed dollar amount.

If your agency doesn't have an electronic system you can also complete form TSP-1 and send it to your agency's payroll or benefits office.

**II. CHOOSE THE AMOUNT OF YOUR CONTRIBUTIONS**

Your choice will cancel all previous elections.

To start or change the amount of traditional (pre-tax) or Roth (after-tax) contributions to your TSP account, enter **either** a whole percentage of your basic pay per pay period **or** a whole dollar amount per pay period for each type of contribution you elect. (You may choose a percentage for one type of contribution and a dollar amount for the other type of contribution.) **Remember:** A blank line next to a type of contribution equals 0% or \$0 contributed.

6. Traditional (Pre-Tax) Contributions	<input style="width: 100%;" type="text" value=""/>	.0%	OR	7. \$ <input style="width: 100%;" type="text" value=""/>	.00
8. Roth (After-Tax) Contributions	<input style="width: 100%;" type="text" value=""/>	.0%	OR	9. \$ <input style="width: 100%;" type="text" value=""/>	.00

In Section II, enter 10% in Box 6 or Box 8 on the TSP-1.

Call TSP at 1-877-968-3778 and choose option 3 for help, or visit the TSP Website, <https://www.tsp.gov/forms/index.html> (Select TSP-1) it includes a short video

[Report Issue](#)

## Appendix F Ballpark Tool Formulas

Total monthly retirement income in today's dollars comprises three parts:

$$\begin{aligned}
 Total &= \underbrace{Annuities}_I + \underbrace{TSP\ Balance}_II + \underbrace{Other\ Income}_III \\
 &= \frac{FERS}{(1+i)^b} + SS + \frac{TotBal + WorkRet + AddPens}{12 \times RAV(R)}
 \end{aligned}$$

### Variable I: Inflation-indexed annuities (SS, FERS)

This variable captures the value of Federal Employee Retirement System(FERS) and Social Security (SS) payments.

- SS = Value of initial monthly SS payment in today's dollars (user input or calculated).
- FERS = Value of initial monthly FERS payment in future dollars (at retirement age) (calculated).

### Variable II: TSP Balance

#### *User Inputs:*

- *CurrTSPBal* = Value of TSP balance in today's dollars
- *ContribRate* = Contribution rate as percent of salary
- *ContribAmt* = Contribution amount in dollars per pay period
- *ContribRateEquiv* = Contribution rate equivalent as percent of salary
- *Salary* = Annual salary in today's dollars
- *OthSav* = Value of additional retirement savings in today's dollars

#### *Calculated Values:*

- *ContribRateEquiv* = Contribution rate equivalent as percent of salary. See equation 2.

$$ContribRateEquiv = \begin{cases} ContribRate, & \text{if rate selected} \\ ContribAmt/Salary, & \text{if amount selected} \end{cases} \quad (2)$$

- $AddlTSPBal$  = Value of future TSP contributions accumulated at  $r_b$  in future dollars (at retirement age). See equation 3.

$$AddlTSPBal = \frac{ContribRateEquiv \times Salary}{n} \times \left[ \left(1 + \frac{g}{n}\right)^{bn} \frac{\left(\frac{1+(r_b/n)}{1+(g/n)}\right)^{bn} - 1}{\frac{1+(r_b/n)}{1+(g/n)} - 1} \right] \quad (3)$$

where  $g$  is annual wage growth,  $n$  is the number of pay periods in a year,  $b$  is the number of years before retirement, and  $r_b$  is expected rate of return before retirement.

- $TotBalAtRet$  = Value of total retirement savings balance in future dollars (at retirement age). See equation 4.

$$TotBalAtRet = (CurrTSPBal + OthSav) \times (1 + r_b)^b + AddTSPBal \quad (4)$$

where  $b$  is the number of years before retirement, and  $r_b$  is expected rate of return before retirement.

- $TotBal$  = Value of total retirement savings balance in today's dollars.

$$TotBal = \frac{TotBalAtRet}{(1 + i)^b} \quad (2)$$

where  $b$  is the number of years before retirement, and  $i$  is the inflation rate.

### Variables III: Other Income

#### *User Inputs:*

- $w$  = Expected income from working after retirement (assumed to be in today's dollars and annual)
- $p$  = Expected pension in retirement (assumed to be in today's dollars and annual amount)

#### *Calculated Values:*

- $WorkRet$  = Value of income from working in retirement in future dollars (at retirement age).

See equation 3

$$WorkRet = w \times \frac{(1 + r_a)^{w_a - 1}}{r_a} \quad (3)$$

where  $r_a$  is the expected rate of return after retirement, and  $w_a$  is the number of years working after retirement

- $AddPens$  = Value of income from additional pension in future dollars (at retirement age).

See equation 4

$$AddPens = p \times \frac{[RAV(R + o)]}{(1 + r_a)^o} \quad (4)$$

where  $RAV(X)$  is the real annuity value function <sup>11</sup>  $R$  is retirement age ,  $o$  is years in retirement before other pension begin, and  $r_a$  is the expected rate of return after retirement.

## Total Monthly Retirement Income Formula

Total monthly retirement income in today's dollars is then calculated as follows:

$$Total = \frac{FERS}{(1 + i)^b} + SS + \frac{TotBal + WorkRet + AddPens}{12 \times RAV(R)} \quad (5)$$

## Definition Glossary

- $i$ : Inflation rate
- $R$ : Retirement age
- $a$ : Number of expected years in retirement
- $b$ : Number of years before retirement
- $o$ : Years in retirement before other pension begins
- $r_b$ : Expected rate of return before retirement
- $r_a$ : Expected rate of return after retirement
- $w_a$ : Number of years working after retirement

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<sup>11</sup>  $RAV(X)$  calculates the value of \$1 paid annually for someone currently age  $x$  until death, growing with inflation each year, valued in today's dollars.



- $n$ : Number of pay periods in a year
- $g$ : Annual wage growth
- $RealAnnuityValue(x)$ : Value of \$1 paid annually for someone currently age  $x$  until death, growing with inflation each year, valued in today's dollars